

周课 7

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1. Non-Parametric Model

- Penalized likelihood;
- smoothing;
- fitting wiggly lines through points;
- semi-parametric models;

- splines

$$Y_i \sim \pi(\lambda_i, \theta)$$

$$g(\lambda_i) = X_i\beta + f(W_i)$$

Where, • Y_i : response

- $\pi(\lambda_i, \theta)$ is the response distribution
- X_i, W_i are covariates
- $f(w)$: is the smoothing function
- $g(\lambda)$: is the link function
- β : coefficients

1. Penalized Likelihood

$$L_P(\beta, f, \alpha; Y) = \log(\pi(Y; \beta, f)) - \alpha \int \left[\frac{\partial^2 f(w)}{\partial w^2} \right]^2 du$$

Where, • α : penalty parameter, $\alpha \uparrow \Rightarrow$ smoother f

- smoother $f(x) \Rightarrow$ smaller $f''(x)$

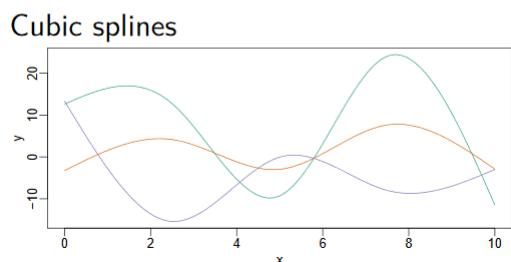
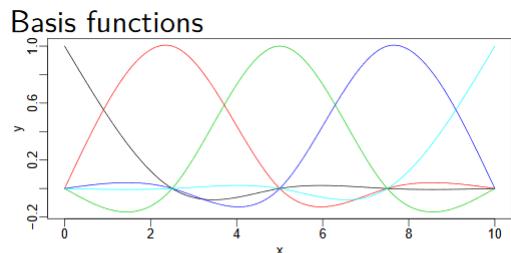
$$\Rightarrow \hat{\beta}(\alpha), \hat{f}(\alpha) = \arg \min_{\beta, f} L_P(\beta, f, \alpha; Y)$$

- A good \hat{f} is a compromise between fitting the data and being smooth.

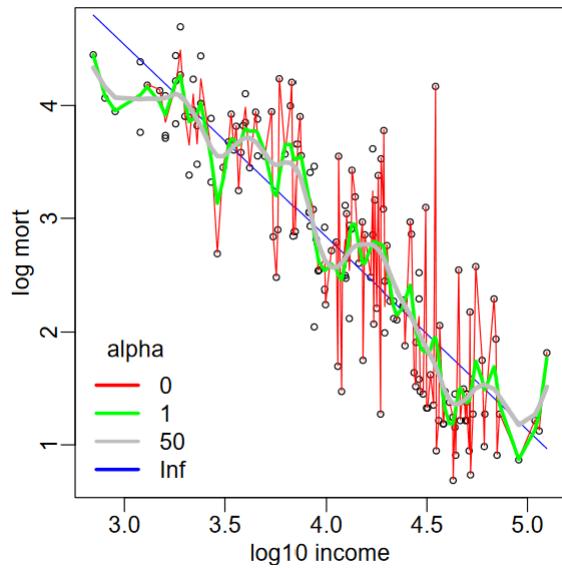
1.2 Cubic Spline

- The f that maximizes the penalized likelihood must be a cubic spline polynomial...

```
knitr:::include_graphics("1.png")
```



```
knitr:::include_graphics("2.png")
```



- The basis function of cubic splines: $ax^3 + bx^2 + cx + d\dots$

Maximizing likelihood over all possible f :

- The larger the α , the smoother the curve (f)...
- When $\alpha \rightarrow \infty$, f is a straight line.

How?:

- Divide your data (evenly) into K subsets, and fit a cubic spline on each subset. Make sure the f function is continuous/1st-order-diff/2nd-order-diff at each knot...

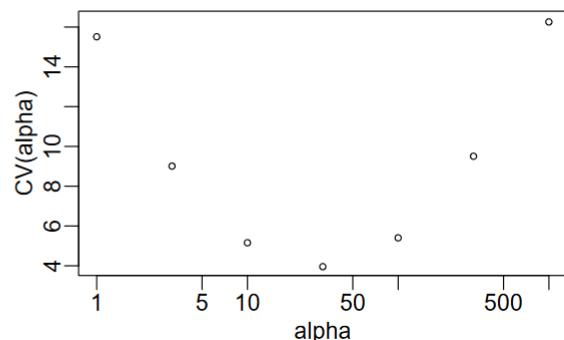
Choosing α : Cross Validation

```
knitr:::include_graphics("3.png")
```

Cross validation

- Find $\hat{\lambda}^{(-i)}$ by excluding observation i
- compute $pr(Y_i|\hat{\lambda}^{(-i)})$
- repeat for $i = 1 \dots N$
- $CV(\alpha) = -\sum_i \log[pr(Y_i|\hat{\lambda}^{(-i)})]$

$$\hat{\alpha} = \operatorname{argmax}_{\alpha} CV(\alpha)$$



2. Generalized Additive Model (GAM)

- Fit a GAM for the Math score data...

```
library('mgcv')

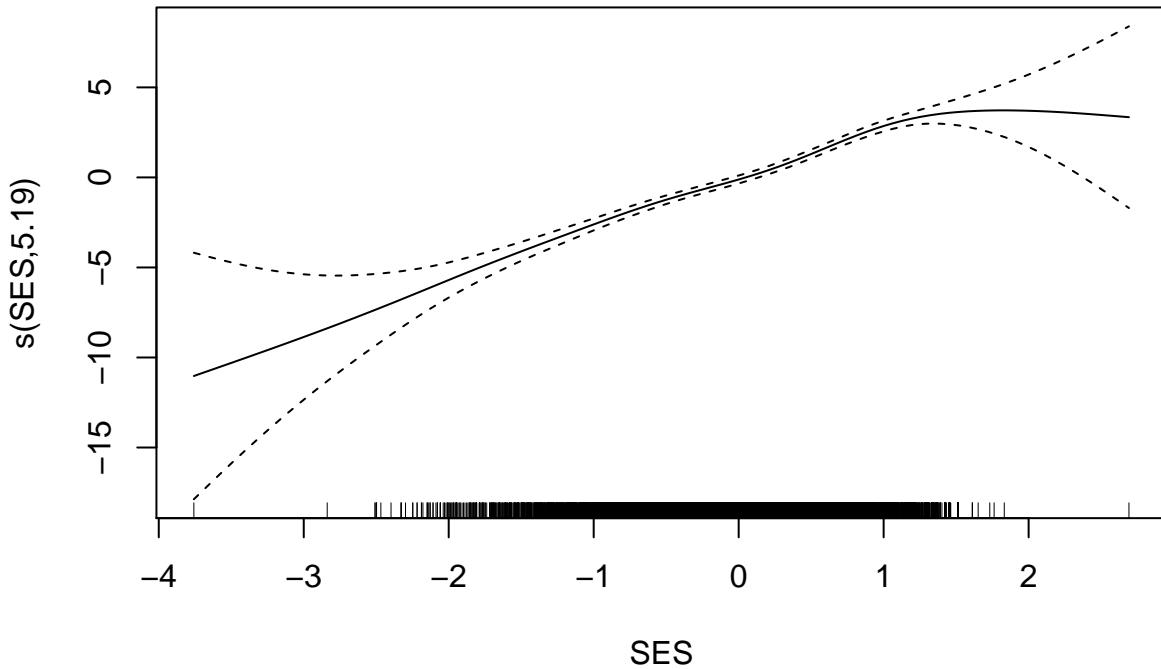
## Loading required package: nlme

## This is mgcv 1.8-28. For overview type 'help("mgcv-package")'.

mathGam =gam(MathAch~s(SES)+Minority*Sex,
             data=MathAchieve)
knitr::kable(summary(mathGam)$p.table[,1:2],
             digits=1)
```

	Estimate	Std. Error
(Intercept)	14.3	0.1
MinorityYes	-2.9	0.2
SexFemale	-1.4	0.2
MinorityYes:SexFemale	0.2	0.3

```
plot(mathGam)
```



```
mathGam$sp # smoothing parameter
```

```
##      s(SES)
## 0.8254378
```

2.1 Smoothing Interation

- Now we fit another GAM, with interaction between the covariates that are being smoothed...

```
mathGamInt =gam(MathAch~s(SES,by=Minority)
+Minority*Sex,
data=MathAchieve)

knitr::kable(summary(mathGamInt)$p.table[,1:2],
digits=1)
```

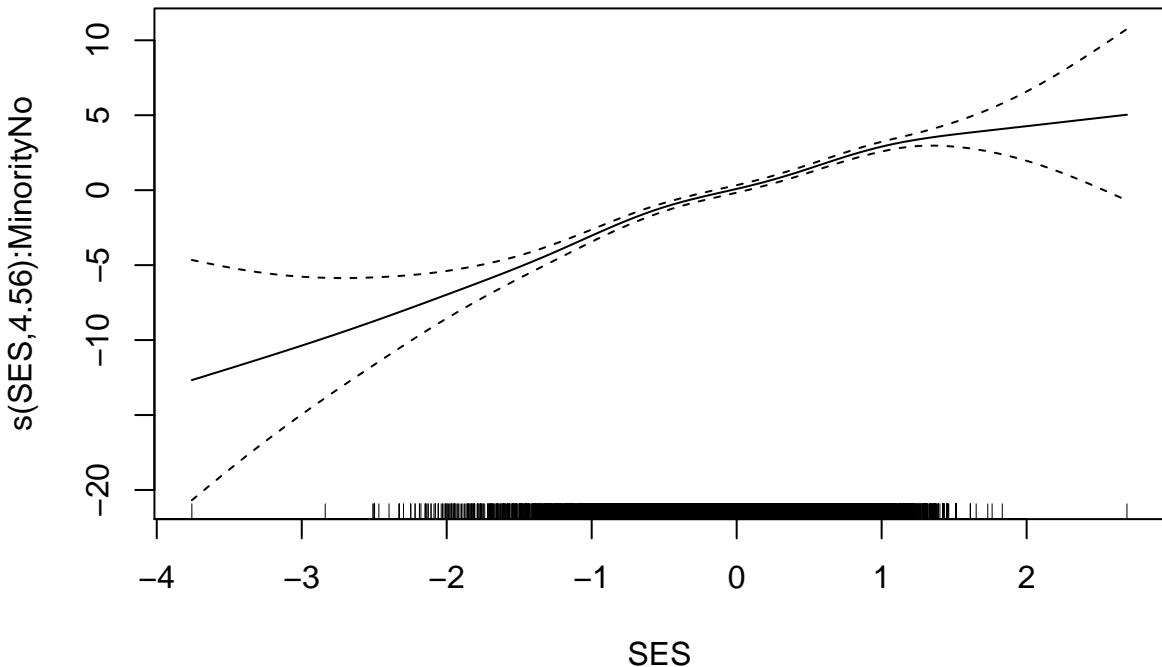
	Estimate	Std. Error
(Intercept)	14.2	0.1
MinorityYes	-3.0	0.3
SexFemale	-1.4	0.2

	Estimate	Std. Error
MinorityYes:SexFemale	0.1	0.3

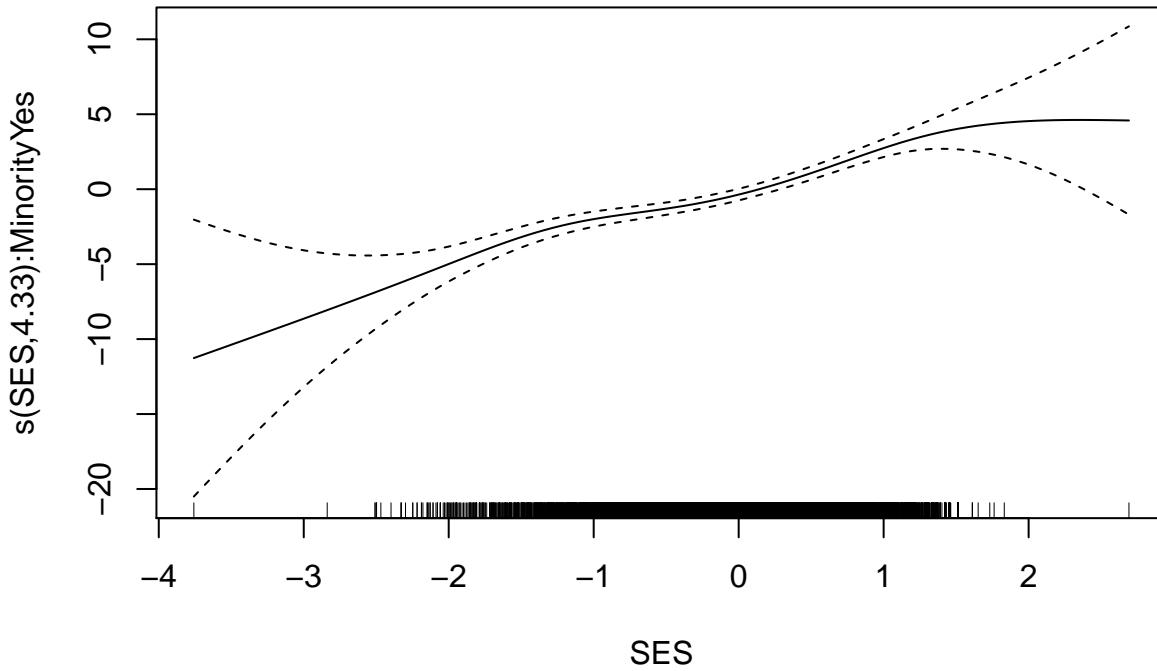
```
mathGamInt$sp
```

```
## s(SES):MinorityNo s(SES):MinorityYes
##          0.820158           0.614983
```

```
# plot the SES/minority
plot(mathGamInt,select =1)
```



```
plot(mathGamInt,select =2)
```



2.2 Common smoothing parameter

```
knitr::include_graphics("4.png")
```

$$\begin{aligned} Y_{ij} &\sim N(\lambda_{ij}, \tau^2) \\ \lambda_{ij} &= X_{ij}\beta + f_i(W_{ij}; \nu) \end{aligned}$$

- Y_{ij} is the observation for individual j in ethnic group i
- X_{ij} is a vector of covariates (ethnic group, sex, interaction)
- $f_i(w; \nu)$ is the smoothly-varying function of SES
 - for ethnic group i
 - with roughness parameter ν .

```
mathGamIntC = gam(MathAch ~ s(SES, by=Minority, id=1) + Minority*Sex,
                  data=MathAchieve)
```

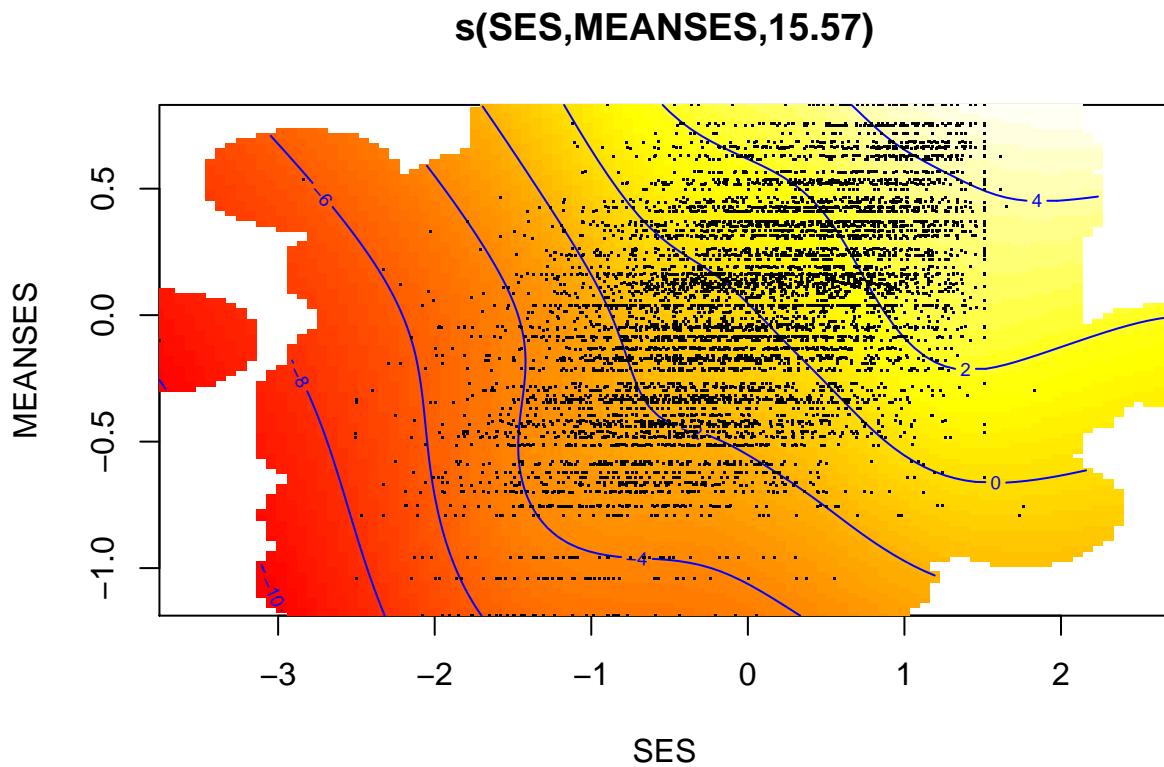
```
mathGamIntC$sp
```

```
## s(SES):MinorityNo
```

```
##          0.7492505
```

2.3 2-D smoothing

```
mathGam2 =gam(MathAch~s(SES, MEANSES)+Minority*+Sex,
               data=MathAchieve)
plot(mathGam2,scheme =2,n2 =100)
```



- If you are from upper class, your score is still likely higher even if your school is weaker...

2.4 Poisson GAM: Ontario deaths

$$Y_i \sim Poisson(\lambda_i)$$

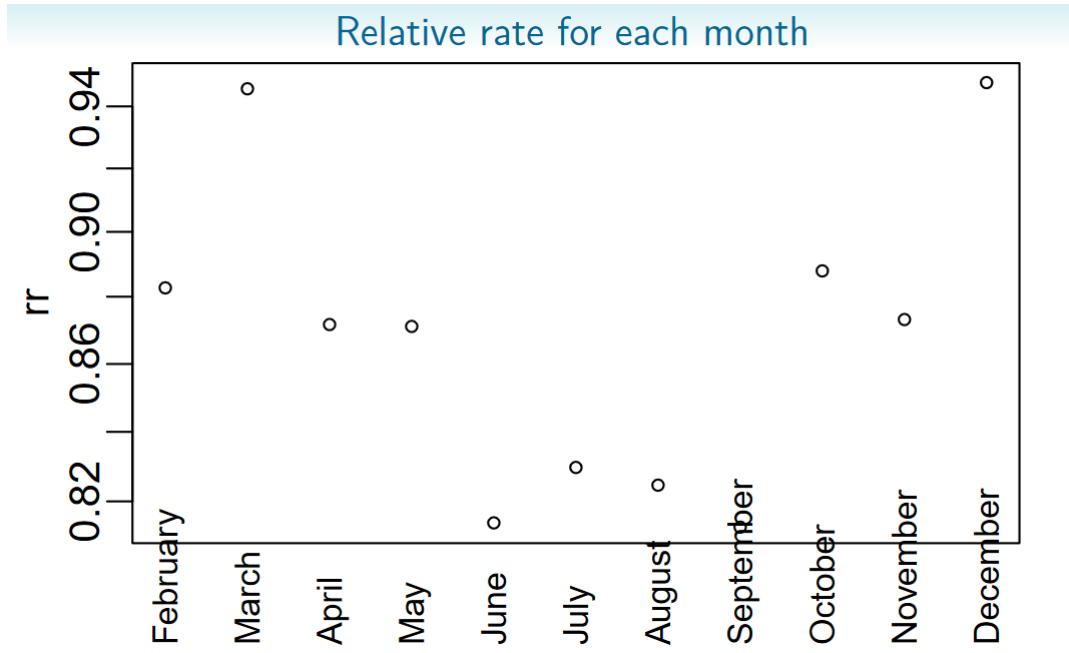
$$\log(\lambda_i) = X_i\beta + f(time)$$

where, • λ_i is the relative rate of death in the i th month

```
deathsGam =gam(~Value~month+s(timeNumeric),
               data=oDeaths,family='poisson'+)
```

```
knitr::kable(summary(deathsGam)$p.table[,1:2],
             digits=3,col.names=c('est','se'))
```

	est	se
(Intercept)	9.001	0.002
monthFebruary	-0.124	0.003
monthMarch	-0.055	0.003
monthApril	-0.137	0.003
monthMay	-0.138	0.003
monthJune	-0.205	0.003
monthJuly	-0.186	0.003
monthAugust	-0.192	0.003
monthSeptember	-0.207	0.003
monthOctober	-0.118	0.003
monthNovember	-0.135	0.003
monthDecember	-0.053	0.003



- relative rate: relative to the baseline, i.e. January...
- Note that different month has different number of days, so, offset!!

$$Y_i \sim Poisson(O_i \lambda_i)$$

$$\log(\lambda_i) = X_i \beta + f(time)$$

where,

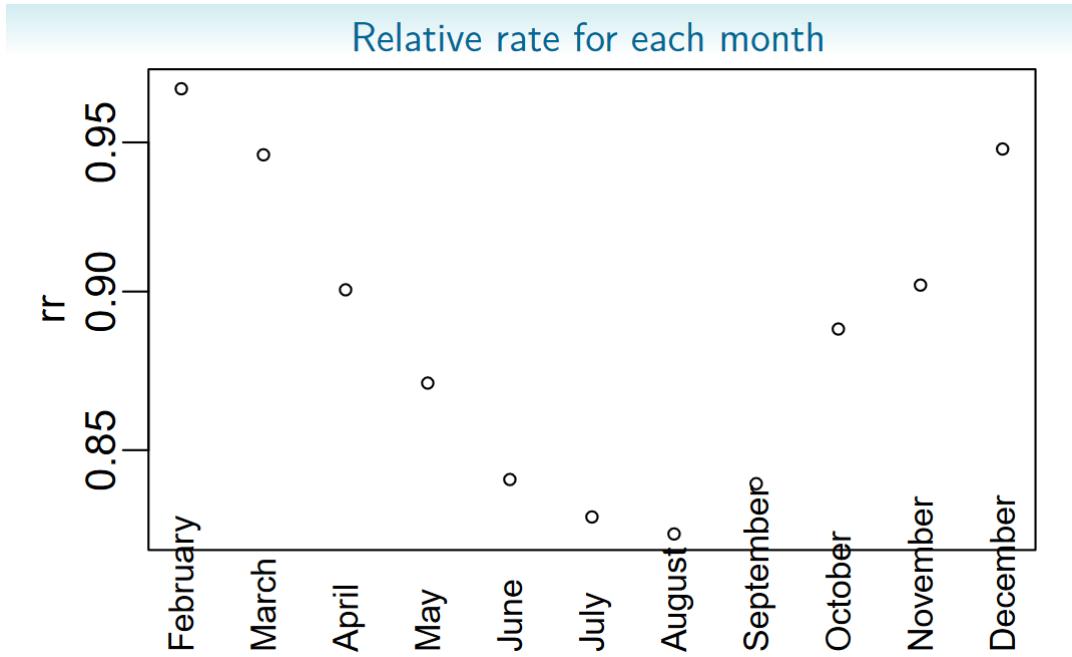
- λ_i is the relative rate of death in the i th month

- O_i is the offset term

```
deathsGam =gam(Value~month+s(timeNumeric)
+offset(nDays),
data=oDeaths,
family='poisson')

knitr::kable(summary(deathsGam)$p.table[,1:2],
digits=3,col.names=c('est','se'))
```

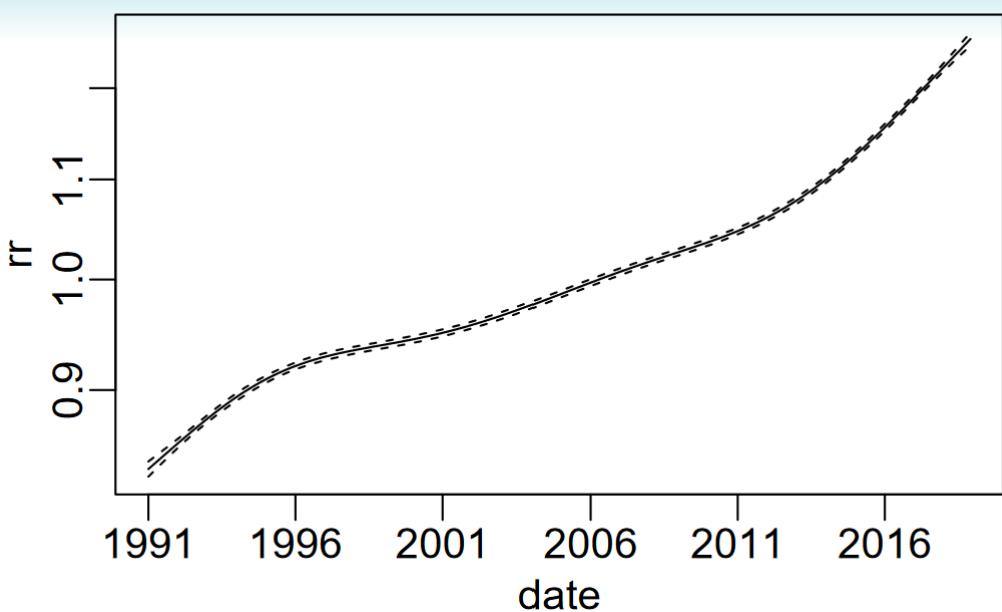
	est	se
(Intercept)	5.567	0.002
monthFebruary	-0.031	0.003
monthMarch	-0.055	0.003
monthApril	-0.104	0.003
monthMay	-0.138	0.003
monthJune	-0.173	0.003
monthJuly	-0.186	0.003
monthAugust	-0.192	0.003
monthSeptember	-0.174	0.003
monthOctober	-0.118	0.003
monthNovember	-0.102	0.003
monthDecember	-0.053	0.003



2.5 Prediction

2.5.1 Trend

```
dSeq =seq(from =min(oDeaths$date),
          by ="5 years",length.out =10)
deathPred =as.matrix(as.data.frame(predict.gam(deathsGam, oDeaths,type ="terms",terms ="s(timeNumeric)))
deathPred =exp(deathPred%*%Pmisc::ciMat())
matplot(oDeaths$timeNumeric, deathPred,log ="y",
        xaxt ="n",xlab ="date",type ="l",
        lty =c(1,+2,2),col ="black",
        ylab ="rr")
axis(1,at =diffftime(dSeq,
                      timeOrigin,units ="days"),
     labels =format(dSeq,"%Y"))
```



2.5.3 Forecasting

```

Stime =seq(from =as.Date("2000/1/1"),
           to =as.Date("2026/1/1"),
           by ="months")

newX =data.frame(timeNumeric
                  =as.numeric(difftime(Stime,
                                       timeOrigin,
                                       units ="days")),
                  month =months(Stime),
                  nDays =log(Hmisc::monthDays(Stime)))

deathsPred =predict(deathsGam, newX, se.fit =TRUE)

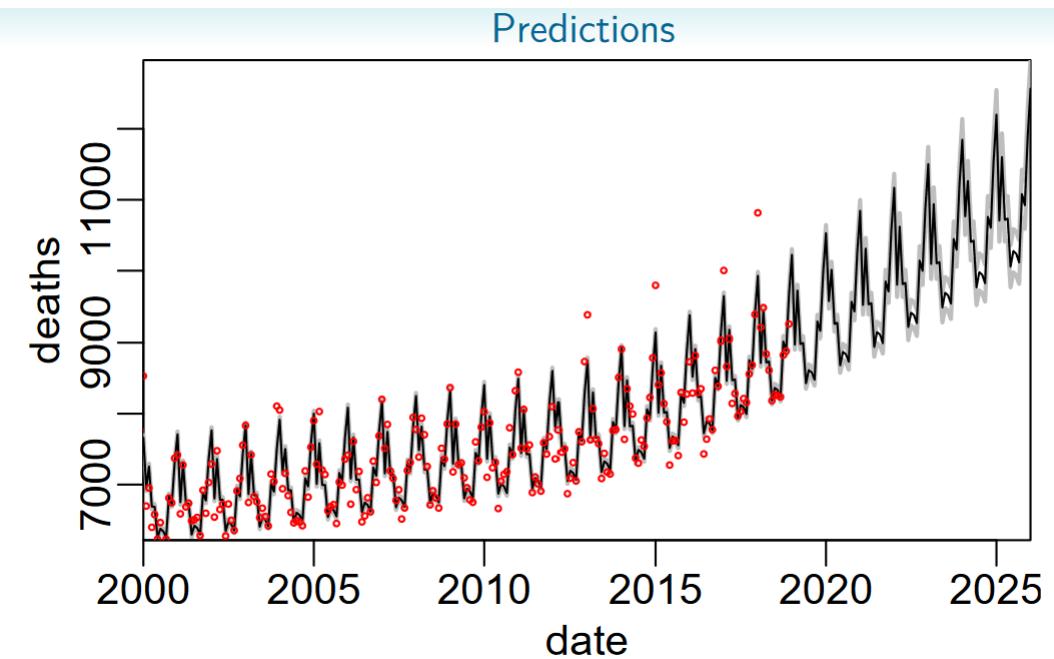
deathsPred =as.data.frame(deathsPred)
deathsPred$lower =deathsPred$fit-2*deathsPred$se.fit
deathsPred$upper =deathsPred$fit+2*deathsPred$se.fit

matplot(Stime,
        exp(deathsPred[,c("lower", "upper", +"fit")]),
        type ="l", lty =1,
        col =c("grey", +"grey", "black"),
        lwd =c(2,2,1), xlab ="date",
        ylab ="deaths", yaxs ="i", xaxs ="i",
        xaxt ="n")

forAxis =seq(from =as.Date("2000/1/1"),

```

```
    to =as.Date("2026/1/1"),
    by ="5 years")
axis(1,as.numeric(forAxis),format(forAxis,"%Y"))
points(oDeaths$date,
       oDeaths$Value,cex =0.5,col ="red")
```



2.6 Change the Parameter constraint

- Add a constant to $f(x)$ doesn't change the penalty;
 - By default, $f(x)$ sums to 0;
 - But we don't know where does $f(x) = 0$;
 - An alternative is to set $f(x_0) = 0$...

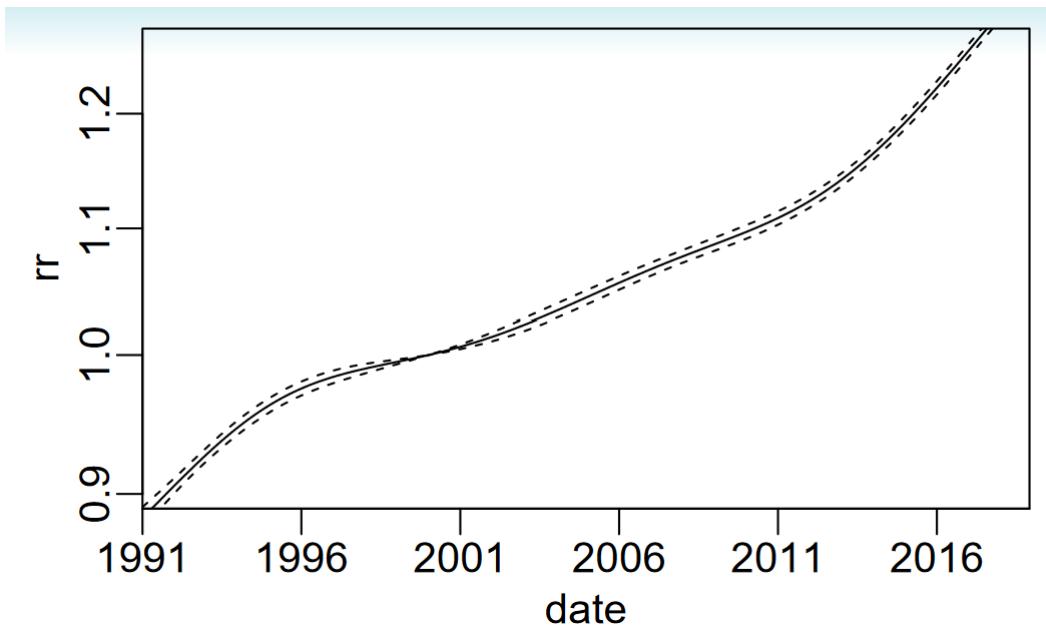
```

            terms ="s(timeNumeric)",
            se.fit =TRUE)))

deathPredC =exp(deathPredC%*%Pmisc::ciMat())
matplot(oDeaths$timeNumeric, deathPredC, log ="y",
        xaxt ="n",xlab ="date",type ="l",
        lty =c(1,+2,2),col ="black",ylab ="rr",
        xaxs ="i",+yaxs ="i",ylim =c(0.89,1.28))
axis(1,at =diffftime(dSeq, timeOrigin,
                      units ="days"),
     labels =format(dSeq,"%Y"))

```

	est	se
(Intercept)	5.510	0.003
monthFebruary	-0.031	0.003
monthMarch	-0.055	0.003
monthApril	-0.104	0.003
monthMay	-0.138	0.003
monthJune	-0.173	0.003
monthJuly	-0.186	0.003
monthAugust	-0.192	0.003
+monthSeptember	-0.174	0.003
monthOctober	-0.118	0.003
monthNovember	-0.102	0.003
monthDecember	-0.053	0.003



2.7 Modelling Seasonality

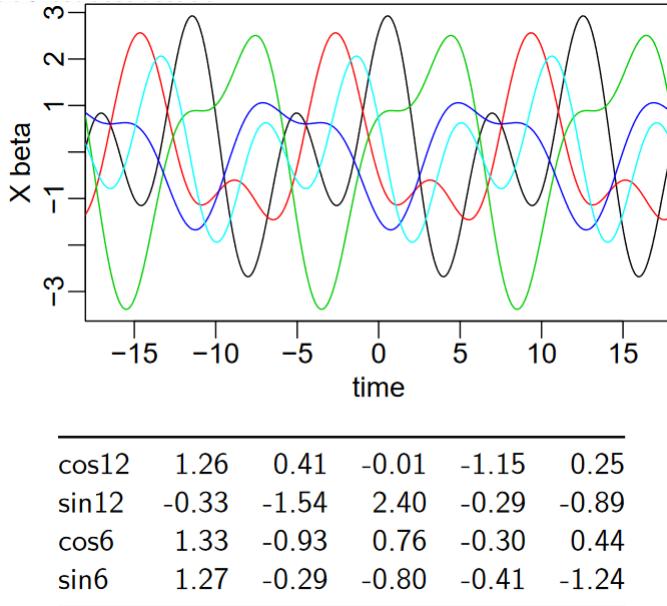
We can see clear seasonality from the month-effect plot.

- The trick to model seasonality is apply trigonometric function as the basis functions...

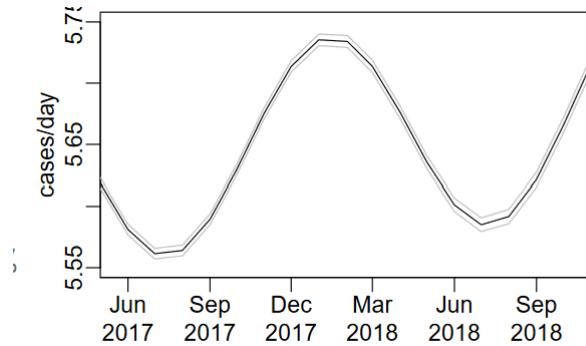
- The monthly effect isn't perfectly sinusoidal
 - use a 12 month and a 6 month frequency

$$\begin{aligned} Y_i &\sim \text{Poisson}(O_i \lambda_i) \\ \log(\lambda_i) &= X_i \beta + f(t_i) \\ X_{i0} &= 1 \end{aligned}$$

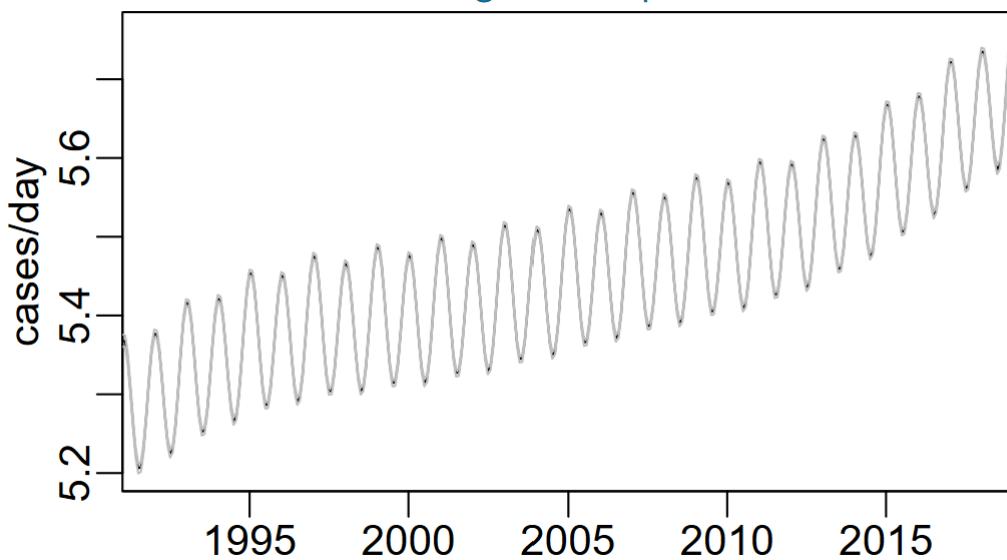
$$\begin{aligned}X_{i1} &= \cos(2\pi t_i/12) \\X_{i2} &= \sin(2\pi t_i/12) \\X_{i3} &= \cos(2\pi t_i/6) \\X_{i4} &= \sin(2\pi t_i/6)\end{aligned}$$



	est	se
(Intercept)	5.456	0.001
cos12	0.085	0.001
sin12	0.017	0.001
cos6	-0.008	0.001
sin6	-0.003	0.001



longer time span

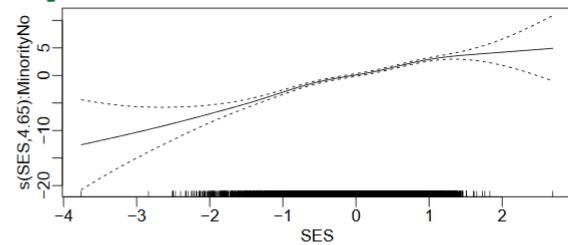


2.8 Number of Knots

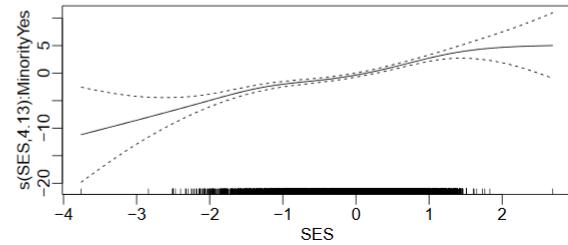
```
> mathGamIntC = gam(MathAch ~
+   s(SES, by=Minority, k=10, id=1) +
+   Minority*Sex,
+   data=MathAchieve)
> mathGamIntC$sp
s(SES):MinorityNo
0.7492505
```

	Estimate	Std. Error
(Intercept)	12.8	0.1
MinorityYes	-2.9	0.2
SexMale	1.4	0.2
MinorityYes:SexMale	-0.1	0.3

```
> plot(mathGamIntC, select = 1)
```

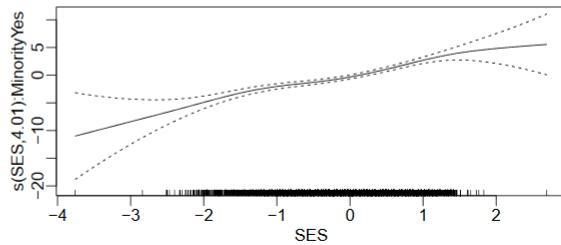
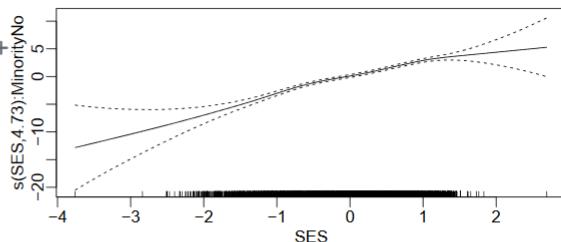


```
> plot(mathGamIntC, select = 2)
```



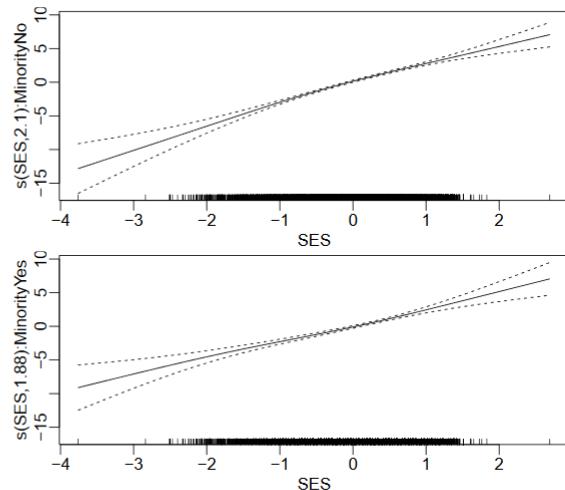
```
> mathGamIntC = gam(MathAch ~
+   s(SES, by=Minority, k=150, id=1) +
+   Minority*Sex,
+   data=MathAchieve)
> mathGamIntC$sp
s(SES):MinorityNo
5105.77
```

	Estimate	Std. Error
(Intercept)	12.8	0.1
MinorityYes	-2.9	0.2
SexMale	1.4	0.2
MinorityYes:SexMale	-0.1	0.3



```
> mathGamIntC = gam(MathAch ~
+   s(SES, by=Minority, k=5, id=1) +
+   Minority*Sex,
+   data=MathAchieve)
> mathGamIntC$sp
s(SES):MinorityNo
4.511766
```

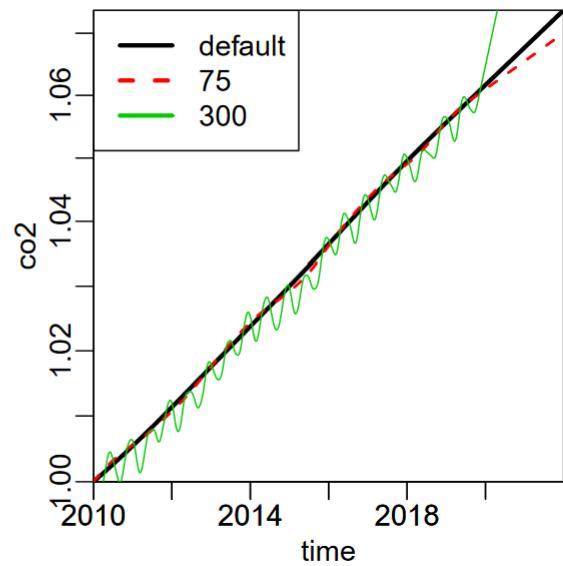
	Estimate	Std. Error
(Intercept)	12.8	0.1
MinorityYes	-2.9	0.2
SexMale	1.4	0.2
MinorityYes:SexMale	-0.1	0.3



- The more knots there are, the better approximation the model is;
- But we don't want that much precision/overfitting...
- If \hat{f} is smooth, we don't need many knots.
- GAM + GCV is fast, but you have to use enough basis functions. (The default number of basis functions is fairly small)

CO2 GAM

```
> res1 = mgcv:::gam(logCo2 ~
+   sin12 + cos12 + sin6 + cos6 +
+   s(timeNumeric, pc=0, k=300),
+   data=co2s)
> res2 = mgcv:::gam(logCo2 ~
+   sin12 + cos12 + sin6 + cos6 +
+   s(timeNumeric, pc=0, k=75),
+   data=co2s)
> resDefault = mgcv:::gam(logCo2 ~
+   sin12 + cos12 + sin6 + cos6 +
+   s(timeNumeric, pc=0), data=co2s)
```



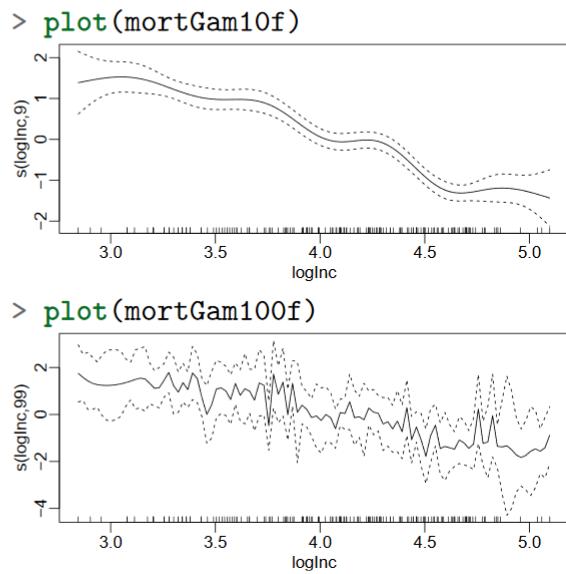
2.8.1 Regression splines

At each subset of the data between 2 knots, do a regression...

Regression splines

- what's faster than GCV?
- don't apply a roughness penalty
- control smoothness with the number of basis functions
- more knots means rougher \hat{f}
- choose knots in some ad-hoc way
- useful for models where GCV isn't possible

```
> mortGam10f = gam(logMort ~
+   s(logInc, k=10, fx=TRUE),
+   data=iMort)
> mortGam100f = gam(logMort ~
+   s(logInc, k=100, fx=TRUE),
+   data=iMort)
```



2.9 ML/model-based smoothing

Use random effects instead of penalized likelihood.

2.9.1 Random Effects

An aside: Random Walks

- RW(0), independent

$$V_t \sim \text{iid } N(0, \tau^2)$$

- RW(1), Brownian motion

$$V_{t+1}|V_k, k < t \sim N(V_t, \tau^2)$$

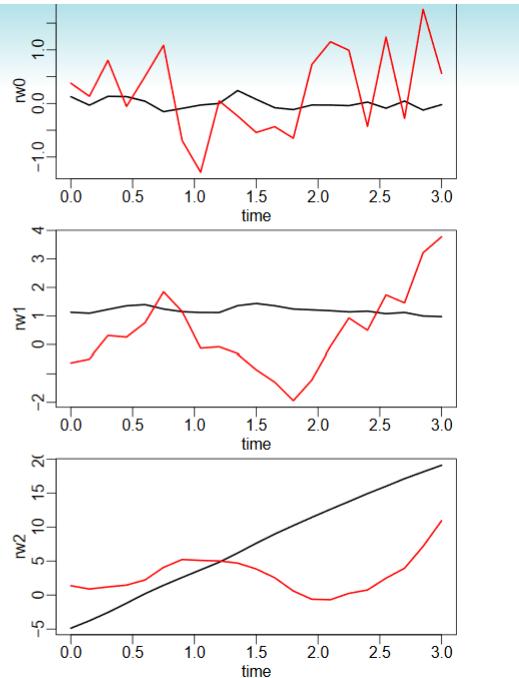
$$V_{t+1} - V_t \sim N(0, \tau^2)$$

- RW(2), Random slope

$$V_{t+1}|V_k, k < t \sim N(-2V_t + V_{t-1}, \tau^2)$$

$$(V_{t+1} - V_t) - (V_t - V_{t-1}) \sim N(0, \tau^2)$$

$$V_{t+1} - 2V_t + V_{t-1} \sim N(0, \tau^2)$$



Using model-based GAM, we can apply the usual statistical machinery:

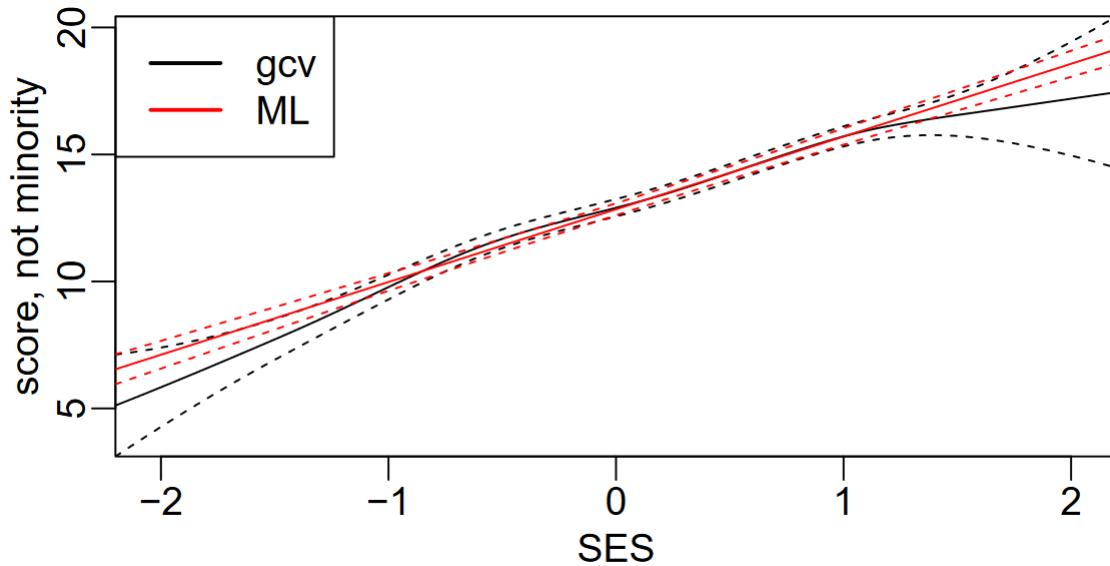
- MLE;

- conditional distribution;
- LR test

2.9.2 ML vs GCV

```
mathGamMl =gam(MathAch~s(SES,by =Minority,
                         id =1,k =30)
                + Minority*Sex,
                data =MathAchieve,
                method ="ML")

mathGamMl$sp
s(SES):MinorityNo
145931.6
```



2.9.3 LR test

Test whether the model above is simply a linear model.

```
mathLm =gam(MathAch~SES*Minority+Minority*Sex,
             data=MathAchieve)

logLik(mathLm,REML =FALSE)
'log Lik.' -23371.28 (df=7)

logLik(mathGamMl)
'log Lik.' -23371.27 (df=7.001989)
```

```
nadiv::LRTest(logLik(mathLm, REML =FALSE),
               logLik(mathGamM1),
               boundaryCorrection =TRUE)$Pval
'log Lik.' 0.5 (df=7)
```

- p-value is pretty large, so, it is sufficient to use a simple linear model for the math data.

2.9.4 Generalized Additive Mixed Model

- GAM's are already GLMM's. So, GAMM is to add additional random effects.

```
mathGamm =gamm4::gamm4(MathAch~SES, k =30) +
           Minority*Sex,
           random =~(1|School),
           data =MathAchieve,
           REML =FALSE)
```

$$\begin{aligned}
 Y_{ij} &\sim N(\lambda_{ij}, \tau^2) \\
 g(\lambda_{ij}) &= X_{ij}\beta + f(W_{ij}) + U_i \\
 [U_1, \dots, U_M]^T &\sim MVN(0, \sigma_1^2 I) \\
 [f(w_1), \dots, f(w_M)]^T &\sim ARIMA_{0,2,1}(\sigma_2^2, 2 - \sqrt{3}) \\
 \text{Or, } f(w) &\sim RW(2) \text{ with variance } \sigma_2^2
 \end{aligned}$$